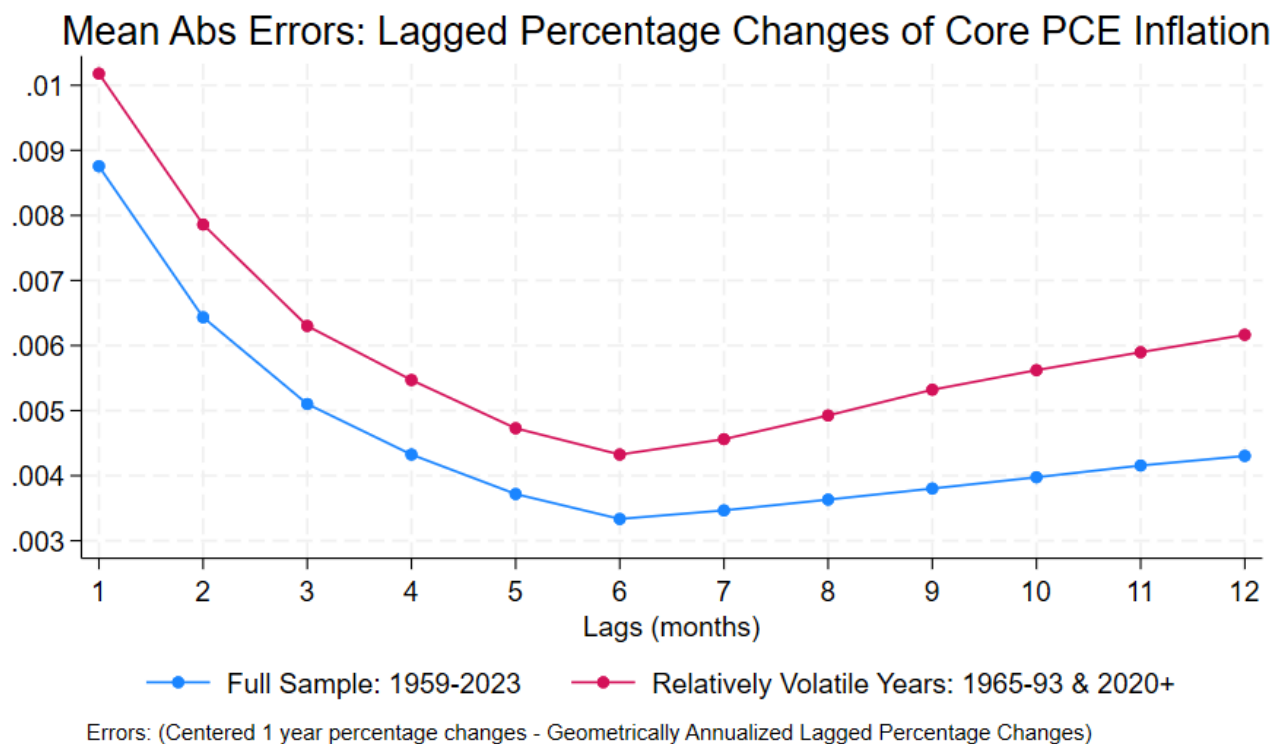
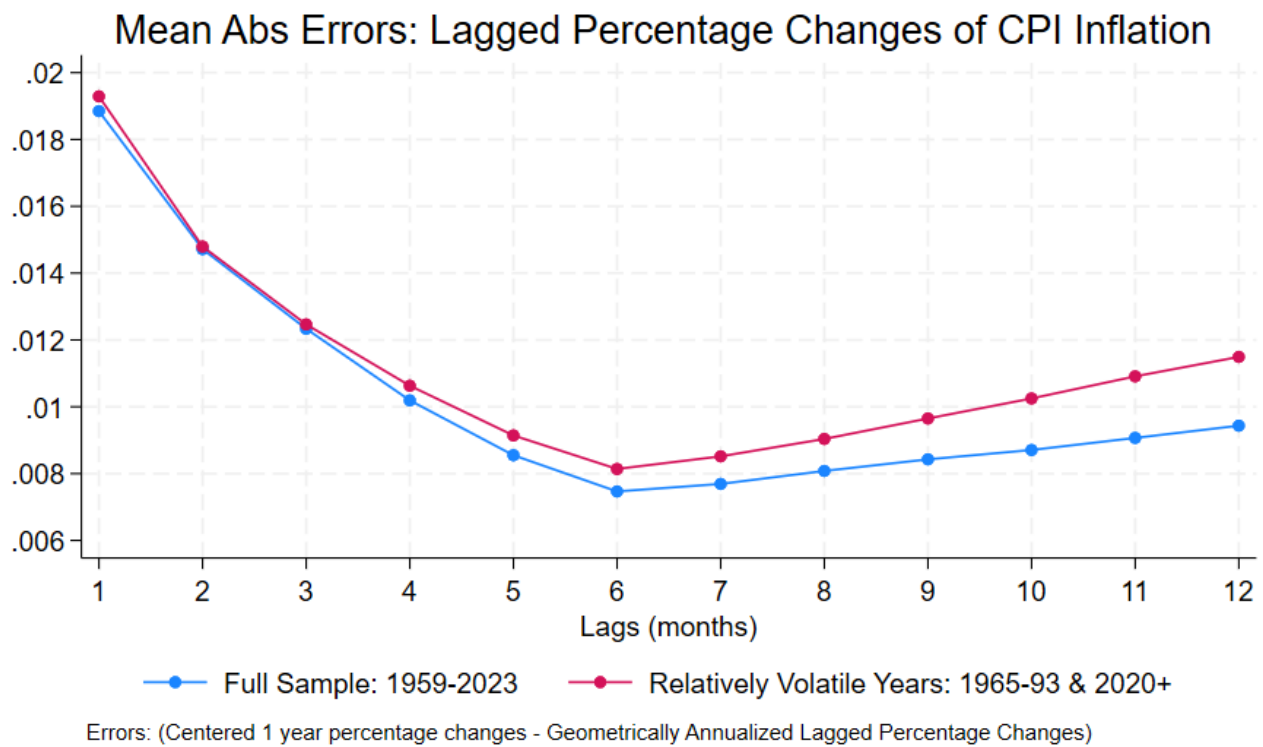
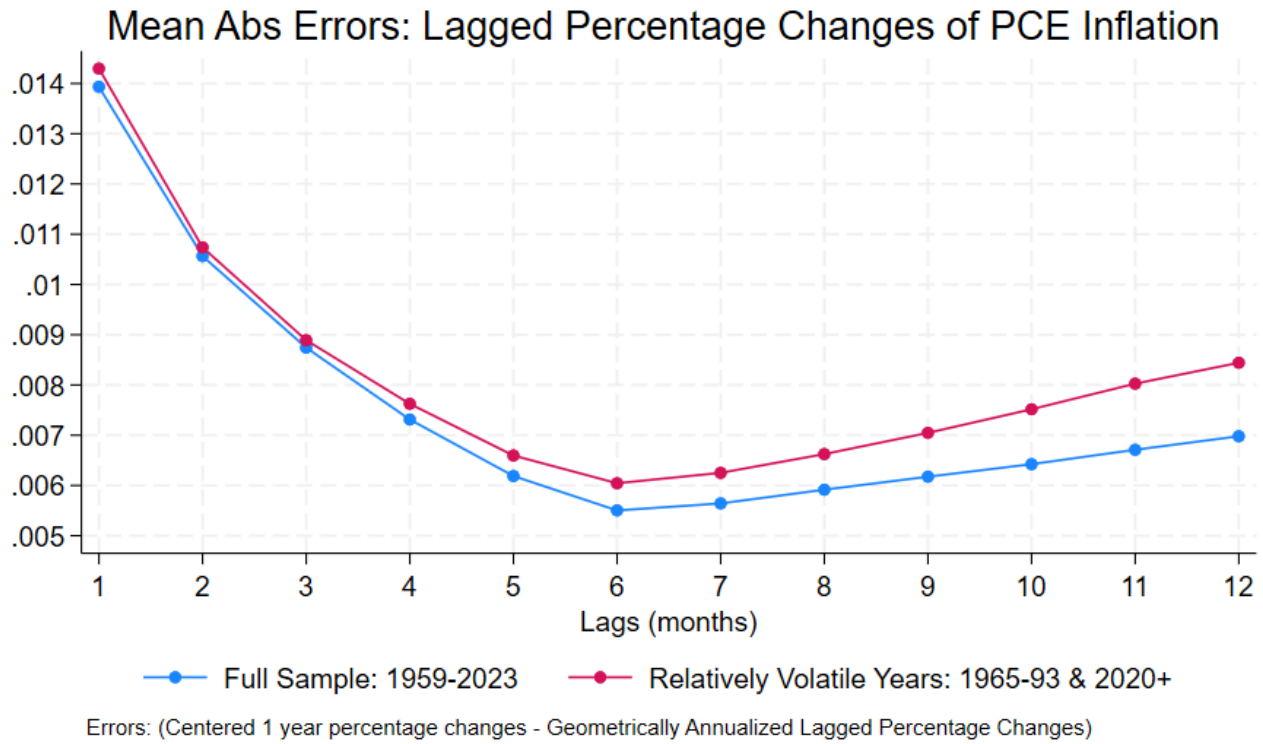
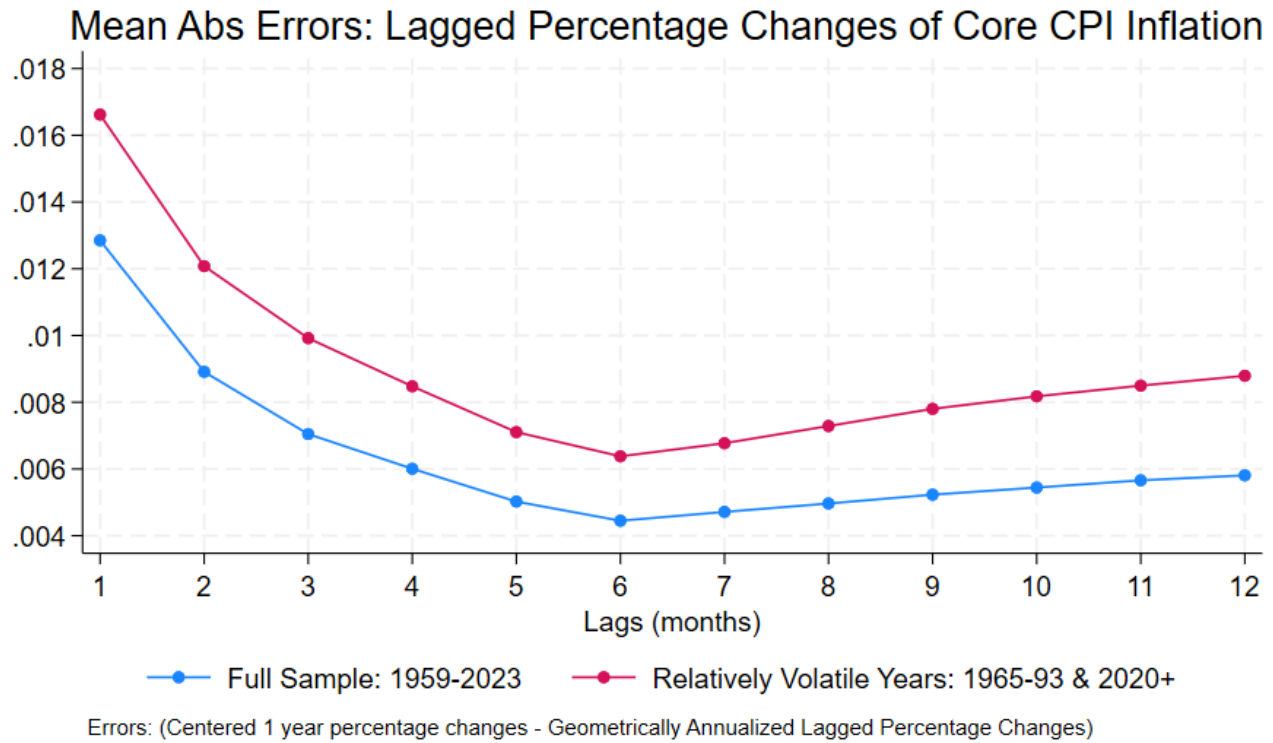


Section 1

Using geometrically annualized returns, 6 month lags fit best for Core PCE, as well as unadjusted PCE, CPI, and Core CPI.







Section 2

Section 3

Table 1 Adjunct

Here I supplement Table 1 with additional summary statistics. The table covers monthly percentage changes from February 1959 through December 2023.

```
. tabstat pCLpcecore pCLpce pCLcpi~core pCLcpi ///
> if sdate>=m(1959m2) & sdate<=m(2023m12) ///
> , format(%8.4f) stat(n mean median sd skew kurtosis q p10 p90 min max)
```

Stats	Core PCE pCLpce~e	PCE pCLpce	Core CPI pCLcpi~e	CPI pCLcpi
N	779.0000	779.0000	779.0000	779.0000
Mean	0.0026	0.0027	0.0030	0.0030
p50	0.0021	0.0022	0.0025	0.0026
SD	0.0020	0.0025	0.0025	0.0031
Skewness	0.9507	0.4875	1.3066	0.2821
Kurtosis	4.4852	5.8555	5.9397	7.1816
p25	0.0013	0.0012	0.0016	0.0012
p50	0.0021	0.0022	0.0025	0.0026
p75	0.0037	0.0039	0.0039	0.0045
p10	0.0007	0.0002	0.0003	0.0000
p90	0.0054	0.0057	0.0062	0.0069

3 Appendix: Forecasting the Present: Optimal Lags for Contemporaneous Core PCE Inflation

Min	-0.0057	-0.0118	-0.0049	-0.0177
Max	0.0103	0.0123	0.0142	0.0181

From MonthlyPrices04b.dta and OptLag24b.do

Comparing p10 and p90 to the median shows that tails are somewhat longer to the right.

Compare Core PCE one month percentage changes with simply annualized and geometrically annualized versions:

```
pcLpcecore: (pcecore/L.pcecore)
pcLapcecore: pcLpcecore*12.
PcLagpcecore: (1+pcLpcecore)^12 -1.
```

```
. tabstat pcLpcecore pcLapcecore pcLagpcecore ///
> if sdate>=m(1959m2) & sdate<=m(2023m12) ///
> , format(%8.4f) stat(n mean median sd skew kurtosis q p10 p90 min max)
```

Stats	pcLpce~e	pcLapc~e	pcLagp~e
N	779.0000	779.0000	779.0000
Mean	0.0026	0.0316	0.0324
p50	0.0021	0.0250	0.0253
SD	0.0020	0.0236	0.0245
Skewness	0.9507	0.9507	1.0335
Kurtosis	4.4852	4.4852	4.6135
p25	0.0013	0.0153	0.0154
p50	0.0021	0.0250	0.0253
p75	0.0037	0.0447	0.0456
p10	0.0007	0.0087	0.0088
p90	0.0054	0.0650	0.0670
Min	-0.0057	-0.0682	-0.0661
Max	0.0103	0.1231	0.1303

Half lives are calculated with the following equation:

Half-life = $\ln(.5)/\ln(\text{ar}(1) \text{ coefficient})$

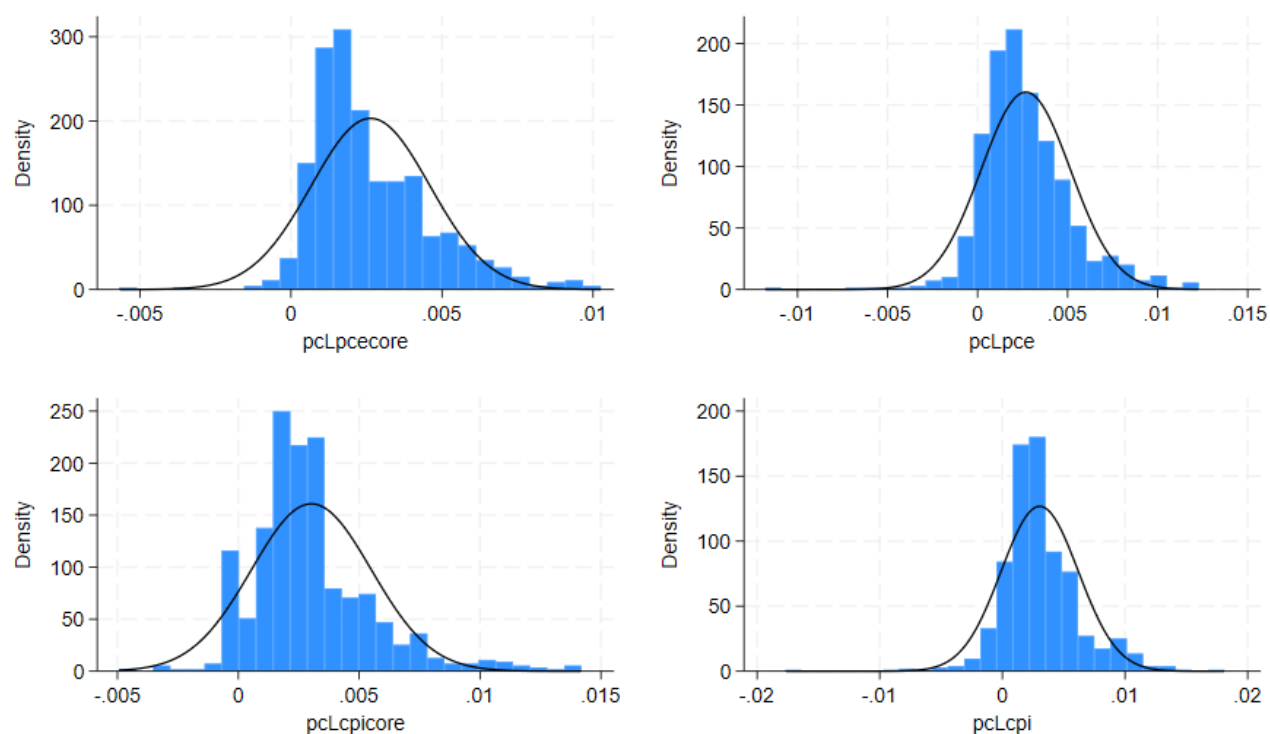
Define the upper bound by taking the ar(1) coefficient and adding 2 standard errors. Define the lower bound by subtracting 2 standard errors.

Half life in Years

	Core PCE	PCE	Core CPI	CPI
Half life, years	3.72	1.13	2.62	0.58
Lower bound	2.00	0.73	1.67	0.41
Upper bound	25.26	2.36	5.93	0.96

The CPI confidence interval is completely outside the Core PCE and Core CPI confidence intervals.

Histogram for monthly percentage changes of core PCE, PCE, core CPI, and CPI.



An AR(1) model is equivalent to an infinite MA(n) model. Solving for n by BIC and AIC metrics gives a rough idea of how far out the influence of a given shock actually persists. Necessarily, smaller sample frames will reduce the power of such tests. Analysis was capped at MA(48). We look at 3 sample frames a) full 1959-2023 sample, b) 1984-2004 Great Moderation sample frame studied by Stock and Watson (2007), and c) 1984-2023 sample, which includes the Great Recession and the COVID supply chain disruption. Lags are stated in months.

	Core PCE	PCE	Core CPI	CPI
MA(n), 1959-2023, min BIC	19	11	11	11
MA(n), 1959-2023, min AIC	32	22	37	22
MA(n), 1984-2004, min BIC	11	1	10	1
MA(n), 1984-2004, min AIC	43	18	47	13
MA(n), 1984-2023, min BIC	9	2	11	2
MA(n), 1984-2023, min AIC	25	22	28	18

The full sample was run using 2 optimization methodologies as a robustness check: they produced the same optimum lags and the same BIC and AIC values for those optimum lags. The BIC lags suggest core CPI and core PCE have memory extending beyond one quarter even during the 1984-2004 era of lower and moderate inflation, while the measures including food and energy had a memory of almost a year during that period. For the full sample Core PCE extended out 19 months by BIC criteria and a full 32 months according to AIC criteria.

Regression results for Core PCE during in the full 1959-2023 sample for MA(19) (lowest BIC) and MA(32) (lowest AIC) follow:

```
. arima dellnpcecore23, ma(1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 ///
> 19) nolog
```

ARIMA regression

```
Sample: 1959m2 thru 2023m12      Number of obs      =      779
                                Wald chi2(19)           =      544.24
Log likelihood = 4194.053          Prob > chi2         =      0.0000
```

		OPG		z	P> z	[95% conf. interval]	
dellnpcecore23	Coefficient	std. err.					
dellnpcecore23							
_cons	.0025959	.0002263	11.47	0.000	.0021524	.0030394	
ARMA							
ma							
L1.	.3491279	.023089	15.12	0.000	.3038743	.3943816	
L2.	.2193244	.0335219	6.54	0.000	.1536227	.2850261	
L3.	.2492281	.0366906	6.79	0.000	.1773158	.3211403	
L4.	.2317146	.0351921	6.58	0.000	.1627394	.3006898	
L5.	.2410805	.0382535	6.30	0.000	.166105	.3160561	
L6.	.3751552	.0406715	9.22	0.000	.2954404	.4548699	
L7.	.3403814	.0422084	8.06	0.000	.2576546	.4231083	
L8.	.2975098	.0413263	7.20	0.000	.2165119	.3785078	
L9.	.3193825	.0425434	7.51	0.000	.235999	.4027659	
L10.	.2641206	.0432189	6.11	0.000	.1794131	.3488281	
L11.	.2182067	.0421566	5.18	0.000	.1355812	.3008321	
L12.	.215335	.0406519	5.30	0.000	.1356588	.2950113	
L13.	.1783892	.0413438	4.31	0.000	.097357	.2594215	
L14.	.0908277	.041311	2.20	0.028	.0098596	.1717959	
L15.	.1506242	.0434093	3.47	0.001	.0655436	.2357048	
L16.	.1883014	.0403326	4.67	0.000	.1092508	.2673519	
L17.	.219713	.0393695	5.58	0.000	.1425502	.2968757	
L18.	.1936183	.0384801	5.03	0.000	.1181988	.2690379	
L19.	.113158	.0386732	2.93	0.003	.0373598	.1889561	
/sigma	.001109	.000019	58.43	0.000	.0010718	.0011462	

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

```
. arima dellnpcecore23, ma(1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18 ///
> 19 20 21 22 23 24 25 26 27 28 29 30 31 32) nolog
```

ARIMA regression

Sample: 1959m2 thru 2023m12

Number of obs = 779

Wald chi2(32) = 443.32

Log likelihood = 4215.098

Prob > chi2 = 0.0000

		OPG				
dellnpcecore23		Coefficient	std. err.	z	P> z	[95% conf. interval]
dellnpcecore23						
	_cons	.0025838	.0003312	7.80	0.000	.0019348 .0032329
ARMA						
	ma					
	L1.	.3159986	.0260459	12.13	0.000	.2649497 .3670476
	L2.	.1981991	.0353528	5.61	0.000	.128909 .2674893
	L3.	.2488131	.0407111	6.11	0.000	.1690208 .3286053
	L4.	.2296119	.03798	6.05	0.000	.1551725 .3040514
	L5.	.2505199	.0417336	6.00	0.000	.1687235 .3323162
	L6.	.3752124	.0417544	8.99	0.000	.2933752 .4570496
	L7.	.3185782	.0439522	7.25	0.000	.2324335 .404723
	L8.	.2898487	.0434359	6.67	0.000	.2047159 .3749815
	L9.	.3159943	.0445484	7.09	0.000	.2286811 .4033075
	L10.	.2775714	.0444467	6.25	0.000	.1904575 .3646852
	L11.	.2587957	.0440156	5.88	0.000	.1725267 .3450648
	L12.	.2945654	.0447143	6.59	0.000	.2069269 .3822039
	L13.	.2824754	.0468471	6.03	0.000	.1906568 .3742941
	L14.	.2157713	.0460768	4.68	0.000	.1254624 .3060803
	L15.	.2733551	.0501143	5.45	0.000	.1751329 .3715773
	L16.	.2979904	.0482954	6.17	0.000	.2033332 .3926476
	L17.	.3252803	.0478865	6.79	0.000	.2314245 .4191362
	L18.	.2955999	.0458932	6.44	0.000	.205651 .3855489
	L19.	.2376097	.048157	4.93	0.000	.1432237 .3319958
	L20.	.1955119	.049909	3.92	0.000	.0976921 .2933318
	L21.	.235883	.0452263	5.22	0.000	.1472411 .324525
	L22.	.1713197	.0502954	3.41	0.001	.0727424 .2698969
	L23.	.1605364	.0450835	3.56	0.000	.0721743 .2488984
	L24.	.1635526	.0460505	3.55	0.000	.0732952 .2538099
	L25.	.1260038	.0430447	2.93	0.003	.0416377 .21037
	L26.	.0966891	.0442868	2.18	0.029	.0098885 .1834896
	L27.	.1330819	.0423015	3.15	0.002	.0501725 .2159914
	L28.	.1036175	.0440008	2.35	0.019	.0173776 .1898575
	L29.	.0852929	.0416842	2.05	0.041	.0035934 .1669924
	L30.	.1147482	.0400216	2.87	0.004	.0363073 .1931892
	L31.	.0928838	.0409749	2.27	0.023	.0125746 .1731931
	L32.	.0766307	.0409195	1.87	0.061	-.0035701 .1568316
	/sigma	.0010792	.0000205	52.55	0.000	.0010389 .0011194

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

For the 1959-2023 specification with the lowest BIC, all MA terms are significant at the 5% level. For the run with the lowest AIC, all MA terms are significant at the 10% level.

Regression results for Core PCE during the 1984-2004 subsample for MA(11) (lowest BIC) and MA(43) (lowest AIC) are reproduced below. In contrast to the full sample, most of the MA terms are not significant at the 5% level:

ARIMA regression

Sample: 1984m1 thru 2004m12

Number of obs = 252

Wald chi2(11) = 66.27

Log likelihood = 1352.249

Prob > chi2 = 0.0000

		OPG		z	P> z	[95% conf. interval]	
		Coefficient	std. err.				
dellnpcecore23							
dellnpcecore23	_cons	.0022025	.0001802	12.22	0.000	.0018493	.0025556
ARMA							
	ma						
	L1.	.0349102	.0446745	0.78	0.435	-.0526501	.1224705
	L2.	-.025183	.0805989	-0.31	0.755	-.183154	.1327881
	L3.	.2212228	.0866598	2.55	0.011	.0513727	.3910728
	L4.	.2871652	.073199	3.92	0.000	.1436979	.4306325
	L5.	.2247761	.083411	2.69	0.007	.0612936	.3882587
	L6.	.2874275	.0912271	3.15	0.002	.1086256	.4662294
	L7.	.0309744	.0805152	0.38	0.700	-.1268324	.1887812
	L8.	.0031232	.0794248	0.04	0.969	-.1525467	.158793
	L9.	.1082586	.0770297	1.41	0.160	-.0427168	.2592341
	L10.	.1049995	.0787399	1.33	0.182	-.0493279	.2593269
	L11.	.1467261	.0787189	1.86	0.062	-.0075601	.3010123
	/sigma	.0011275	.0000312	36.10	0.000	.0010662	.0011887

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

ARIMA regression

Sample: 1984m1 thru 2004m12

Number of obs = 252

Wald chi2(35) = 82.63

Log likelihood = 1383.526

Prob > chi2 = 0.0000

		OPG		z	P> z	[95% conf. interval]	
		Coefficient	std. err.				
dellnpcecore23							
dellnpcecore23	_cons	.002329	.000351	6.64	0.000	.001641	.003017
ARMA							
	ma						
	L1.	-.1283078	.	.	.	.	.
	L2.	-.0803945	.1330308	-0.60	0.546	-.3411302	.1803412
	L3.	.1898184	.	.	.	.	.
	L4.	.2152878	.2484426	0.87	0.386	-.2716509	.7022264

L5.	.0792399	.1506355	0.53	0.599	-.2160002	.37448
L6.	.2138848	.150312	1.42	0.155	-.0807213	.508491
L7.	.0740753	.	.	.	.	.
L8.	.0206695	.1617577	0.13	0.898	-.2963697	.3377086
L9.	.1539423	.1375228	1.12	0.263	-.1155975	.423482
L10.	.0747098	.	.	.	.	.
L11.	.1729866	.	.	.	.	.
L12.	.2415938	.1942855	1.24	0.214	-.1391988	.6223864
L13.	.1556449	.148714	1.05	0.295	-.1358293	.4471191
L14.	-.0060402	.	.	.	.	.
L15.	.1675236	.182318	0.92	0.358	-.1898131	.5248604
L16.	.0740143	.	.	.	.	.
L17.	.2832978	.2058333	1.38	0.169	-.120128	.6867236
L18.	.2526207	.1448565	1.74	0.081	-.0312928	.5365343
L19.	.117857	.1475729	0.80	0.425	-.1713806	.4070947
L20.	.1848281	.1128259	1.64	0.101	-.0363066	.4059629
L21.	.2091923	.2271626	0.92	0.357	-.2360381	.6544227
L22.	-.008247	.1906526	-0.04	0.965	-.3819192	.3654252
L23.	.1797052	.1455835	1.23	0.217	-.1056331	.4650436
L24.	.1803261	.1828336	0.99	0.324	-.1780212	.5386735
L25.	.1288013	.1288545	1.00	0.318	-.1237489	.3813516
L26.	.0933022	.130773	0.71	0.476	-.1630081	.3496125
L27.	.1810078	.1475332	1.23	0.220	-.1081519	.4701675
L28.	.045307	.1640956	0.28	0.782	-.2763144	.3669284
L29.	.0745893	.115192	0.65	0.517	-.1511828	.3003614
L30.	.2871661	.1474179	1.95	0.051	-.0017676	.5760998
L31.	.0486	.	.	.	.	.
L32.	-.0021713	.1140475	-0.02	0.985	-.2257002	.2213576
L33.	.0926933	.133706	0.69	0.488	-.1693656	.3547521
L34.	.1895689	.1225895	1.55	0.122	-.050702	.4298398
L35.	.1598855	.1554903	1.03	0.304	-.1448699	.4646409
L36.	.3818094	.134953	2.83	0.005	.1173063	.6463124
L37.	-.1100965	.1333486	-0.83	0.409	-.371455	.1512621
L38.	.071721	.1323208	0.54	0.588	-.187623	.331065
L39.	.2716365	.1526777	1.78	0.075	-.0276062	.5708793
L40.	.1005482	.1086569	0.93	0.355	-.1124153	.3135118
L41.	.1131648	.1136521	1.00	0.319	-.1095892	.3359188
L42.	-.047768	.0992388	-0.48	0.630	-.2422725	.1467364
L43.	-.2935904	.1310958	-2.24	0.025	-.5505334	-.0366474

/sigma	.0009253	.0001142	8.11	0.000	.0007016	.0011491

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

In the 1984-2004 subsample, the MA(11), 5 out of 11 lags were significant at 10% and 7 of 11 has a z-value greater than one. For MA(43), only 5 of 43 lags were significant at 10%, while 17 of 43 (or 40%) had a z-value greater than one. Nonetheless, even in the MA(43) version, 3 out of the last 10 terms were significant at the 10% level of confidence implying at least a small degree of shock persistence 3 years out. A less reassuring 2 of the last 20 terms were significant at the 5% level of confidence.

Finally, we'll look at the 1984-2023 subsample, which includes instability during the Great Recession and the COVID supply chain disruption. Best BIC for Core PCE was MA(9) while best AIC was MA(25). The results are qualitatively similar to the full sample.

```
. arima dellnpcecore23 if year>=1984 & year<=2023, ma(1 2 3 4 5 6 7 8 9 ) nolog ///
> ///
> iterate(600)
```

ARIMA regression

```
Sample: 1984m1 thru 2023m12      Number of obs   =      480
                                Wald chi2(9)        =     184.65
Log likelihood = 2591.34          Prob > chi2      =     0.0000
```

	OPG					
dellnpcecore23	Coefficient	std. err.	z	P> z	[95% conf. interval]	

dellnpcecore23						
_cons	.0019745	.0001317	14.99	0.000	.0017164	.0022326

ARMA						
ma						
L1.	.2212576	.0301824	7.33	0.000	.1621013	.2804139
L2.	.0925407	.0503102	1.84	0.066	-.0060654	.1911469
L3.	.1532423	.049366	3.10	0.002	.0564867	.2499979
L4.	.1574008	.0449503	3.50	0.000	.0692998	.2455019
L5.	.1519843	.0508968	2.99	0.003	.0522284	.2517401
L6.	.2804258	.0472473	5.94	0.000	.1878227	.3730288
L7.	.1710018	.052138	3.28	0.001	.0688131	.2731905
L8.	.1519439	.0473497	3.21	0.001	.0591402	.2447477
L9.	.1540403	.0452539	3.40	0.001	.0653443	.2427362

/sigma	.0010934	.0000207	52.84	0.000	.0010528	.0011339

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

```
. arima dellnpcecore23 if year>=1984 & year<=2023, ma(1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18
> ///
> 19 20 21 22 23 24 25) nolog ///
> ///
> iterate(600)
```

ARIMA regression

```
Sample: 1984m1 thru 2023m12      Number of obs      =      480
                                Wald chi2(25)          =      178.17
Log likelihood = 2612.691          Prob > chi2       =      0.0000
```

dellnpcecore23	Coefficient	OPG std. err.	z	P> z	[95% conf. interval]	

dellnpcecore23 _cons	.0019797	.0002138	9.26	0.000	.0015607	.0023987

ARMA						
ma						
L1.	.1784406	.0331436	5.38	0.000	.1134803	.2434008
L2.	.0532987	.0555305	0.96	0.337	-.0555391	.1621364
L3.	.1249215	.0513344	2.43	0.015	.0243079	.2255352
L4.	.1600502	.0498552	3.21	0.001	.0623358	.2577646
L5.	.1375174	.0517955	2.66	0.008	.0360001	.2390347
L6.	.2612812	.0533752	4.90	0.000	.1566677	.3658946
L7.	.1559928	.0595889	2.62	0.009	.0392007	.2727848
L8.	.149703	.0542242	2.76	0.006	.0434256	.2559805
L9.	.1797308	.0546776	3.29	0.001	.0725647	.286897
L10.	.131274	.056669	2.32	0.021	.0202048	.2423431
L11.	.1255397	.0556681	2.26	0.024	.0164323	.2346471
L12.	.1249804	.0502941	2.48	0.013	.0264058	.223555
L13.	.134104	.0553611	2.42	0.015	.0255982	.2426098
L14.	.0256646	.0531642	0.48	0.629	-.0785353	.1298646
L15.	.1187308	.0575807	2.06	0.039	.0058746	.231587
L16.	.1398913	.0589595	2.37	0.018	.0243329	.2554497
L17.	.1886208	.0535568	3.52	0.000	.0836513	.2935902
L18.	.1409799	.0618518	2.28	0.023	.0197527	.2622071
L19.	.0900123	.065676	1.37	0.171	-.0387103	.2187348
L20.	.1071035	.0585469	1.83	0.067	-.0076462	.2218532
L21.	.1667592	.0529338	3.15	0.002	.0630108	.2705076
L22.	.0905922	.0597509	1.52	0.129	-.0265174	.2077017
L23.	.0671505	.0547764	1.23	0.220	-.0402092	.1745102
L24.	.0860442	.0547362	1.57	0.116	-.0212369	.1933252
L25.	.0823697	.0526208	1.57	0.118	-.0207652	.1855046

/sigma	.0010447	.0000241	43.36	0.000	.0009975	.0010919

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

For the 1984-2023 subsample, all but one of the MA(9) entries was significant at the 5% level of confidence. For MA(25), 18 of 25 were significant at 10%. For the longest 10 lags, 5 were significant at the 10% level of confidence. Persistent memory does not appear to be a quality limited to 1970s inflation.

Section 4

Section 5

Section 6 Best lags for analyzing changes in inflation

Copy of Figure 2:

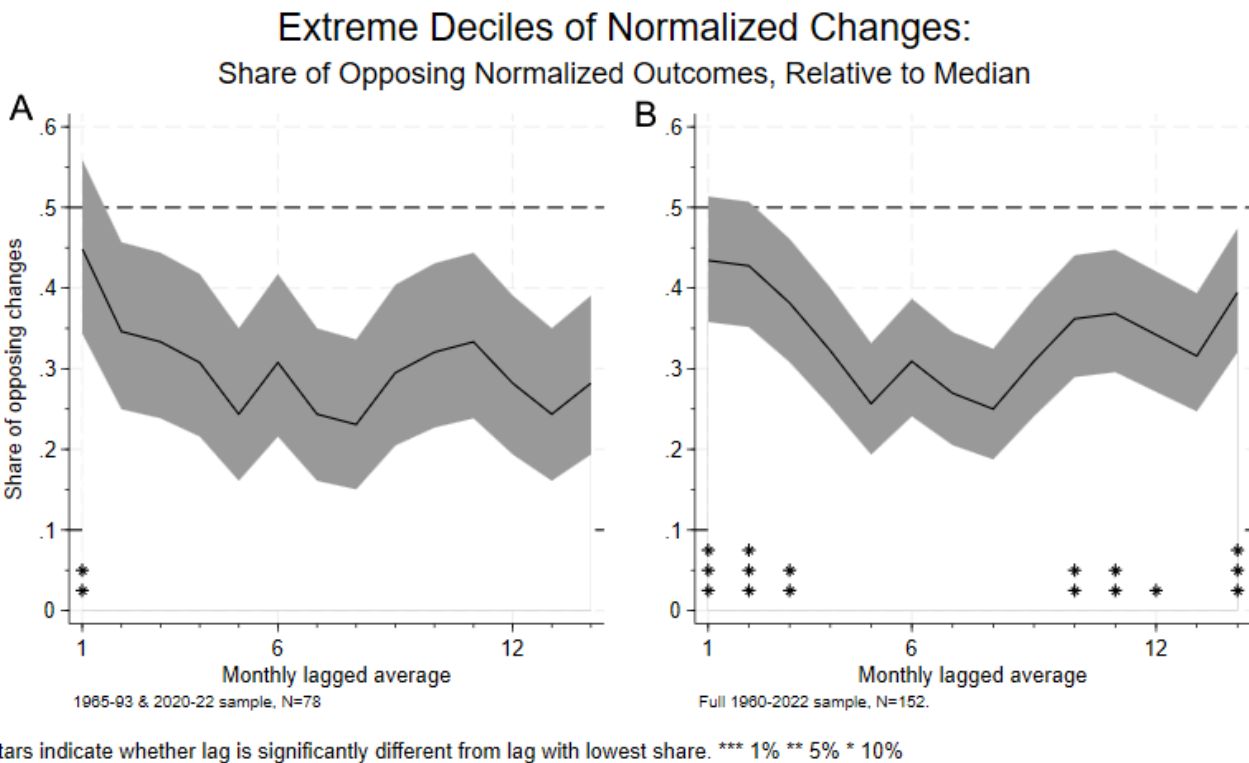


Table for Figure 2, Chart A with confidence intervals based on Agresti and Coull (1998), as recommended by Brown, Cai, and DasGupta (2001)

Proportions for double smoothed share (4), 95% conf interv: Agresti ci

Lag	lower	proportion	upper	N
1	.34331811	.44871795	.55893192	78
2	.24981182	.34615385	.45693828	78
3	.23844238	.33333333	.44387022	78
4	.21593331	.30769231	.41750431	78
5	.16111916	.24358974	.350131	78
6	.21593331	.30769231	.41750431	78
7	.16111916	.24358974	.350131	78
8	.15043294	.23076923	.33637973	78
9	.20479811	.29487179	.40420201	78
10	.22714885	.32051282	.43072626	78
11	.23844238	.33333333	.44387022	78
12	.19374592	.28205128	.39081671	78
13	.16111916	.24358974	.350131	78
14	.19374592	.28205128	.39081671	78
15	.23844238	.33333333	.44387022	78
16	.21593331	.30769231	.41750431	78
17	.21593331	.30769231	.41750431	78
18	.19374592	.28205128	.39081671	78
19	.20479811	.29487179	.40420201	78

20	.24981182	.34615385	.45693828	78
21	.27277121	.37179487	.48285387	78
22	.20479811	.29487179	.40420201	78
23	.22714885	.32051282	.43072626	78
24	.22714885	.32051282	.43072626	78

Confidence intervals based on the Clopper Pearson (1934) exact Binomial are similar:

Proportions for double smoothed share (4), 95% conf interv: Exact ci

Lag	lower	proportion	upper	N
1	.33585694	.44871795	.56559087	78
2	.24196715	.34615385	.46241176	78
3	.23058524	.33333333	.44916669	78
4	.20807964	.30769231	.42242525	78
5	.15347456	.24358974	.3539695	78
6	.20807964	.30769231	.42242525	78
7	.15347456	.24358974	.3539695	78
8	.14287281	.23076923	.33997266	78
9	.19696165	.29487179	.40892378	78
10	.21928851	.32051282	.43583865	78
11	.23058524	.33333333	.44916669	78
12	.18593799	.28205128	.39533118	78
13	.15347456	.24358974	.3539695	78
14	.18593799	.28205128	.39533118	78
15	.23058524	.33333333	.44916669	78
16	.20807964	.30769231	.42242525	78
17	.20807964	.30769231	.42242525	78
18	.18593799	.28205128	.39533118	78
19	.19696165	.29487179	.40892378	78
20	.24196715	.34615385	.46241176	78
21	.26497731	.37179487	.48866119	78
22	.19696165	.29487179	.40892378	78
23	.21928851	.32051282	.43583865	78
24	.21928851	.32051282	.43583865	78

Here are the Agresti and Coull confidence intervals for the full sample presented in Chart B of Figure 2:

Proportions for double smoothed share (4) and conf interv: Agresti ci

Lag	lower	proportion	upper
1	.35798011	.43421053	.51368433
2	.35170034	.42763158	.50713055
3	.30811999	.38157895	.46087602
4	.25310908	.32236842	.40038493
5	.19349252	.25657895	.33166593
6	.24105146	.30921053	.38677543
7	.2052763	.26973684	.34554926
8	.18762887	.25	.32469602
9	.24105146	.30921053	.38677543
10	.28965103	.36184211	.4408443
11	.29579308	.36842105	.44753581
12	.27131245	.34210526	.42068222
13	.24707231	.31578947	.39358813
14	.32050306	.39473684	.47416005
15	.30811999	.38157895	.46087602
16	.26522937	.33552632	.41393175
17	.28965103	.36184211	.4408443
18	.27741054	.34868421	.42741768
19	.28352347	.35526316	.43413831
20	.29579308	.36842105	.44753581
21	.30811999	.38157895	.46087602
22	.24105146	.30921053	.38677543
23	.2591615	.32894737	.40716605
24	.2591615	.32894737	.40716605

Section 7

Nowcasting inflation's direction using data from the relatively volatile inflation era, 1965-1993 and Jan 2020 – Apr 2024.

Motivation:

Investors want to know where inflation is during periods of inflation volatility. But they want to know inflation's direction all the time, because they want to know whether they are entering a period of higher inflation. Moreover, analysis based upon RVIE only (the relatively volatile inflation era) was weak in the center quintile of probabilities, possibly because they were artificially selected out.

So I'll be redoing the analysis with a broader dataset. That means that I'll be taking the full 1960-2019 data to estimate standard deviation and kurtosis to apply to the model. Here are the estimates:

```
sum err1yrCforL6 if RVIE==1 & year<1995, det // So 1965-1993
```

err1yrCforL6				

	Percentiles	Smallest		
1%	-.0164346	-.0202925		
5%	-.0087437	-.0195714		
10%	-.006079	-.0190785	Obs	348
25%	-.0034021	-.0164346	Sum of wgt.	348
50%	-.0001226		Mean	.0000516
		Largest	Std. dev.	.0058413
75%	.0032856	.0196592		
90%	.0064082	.0228871	Variance	.0000341
95%	.0079165	.0230579	Skewness	.3799975
99%	.0196592	.0246152	Kurtosis	5.998675

Contrast with the fuller sample:

```
. dirange pcL6agpcecore pcagC13mpcecore delhat6mn6ga err1yrCforL6
```

```
<SNIP>
```

```
err1yrCforL6 | 559/1330      1959m7 2023m10      0
```

```
. sum err1yrCforL6, det
```

err1yrCforL6				

	Percentiles	Smallest		
1%	-.0120734	-.0202925		
5%	-.0068645	-.0195714		
10%	-.0048139	-.0190785	Obs	772
25%	-.0025424	-.0164346	Sum of wgt.	772
50%	-.0001755		Mean	-6.30e-06
		Largest	Std. dev.	.0047745
75%	.0022934	.0196592		
90%	.0049357	.0228871	Variance	.0000228
95%	.0072557	.0230579	Skewness	.4394425
99%	.0143681	.0246152	Kurtosis	7.143043

For the full sample the standard deviation is lower, but the kurtosis is higher. I think I can live with that.

But here, this Appendix will consider the RVIE subsample, walking through the results.

Reference: Stata Code for Student's t:

```
* rv21a: Calc:

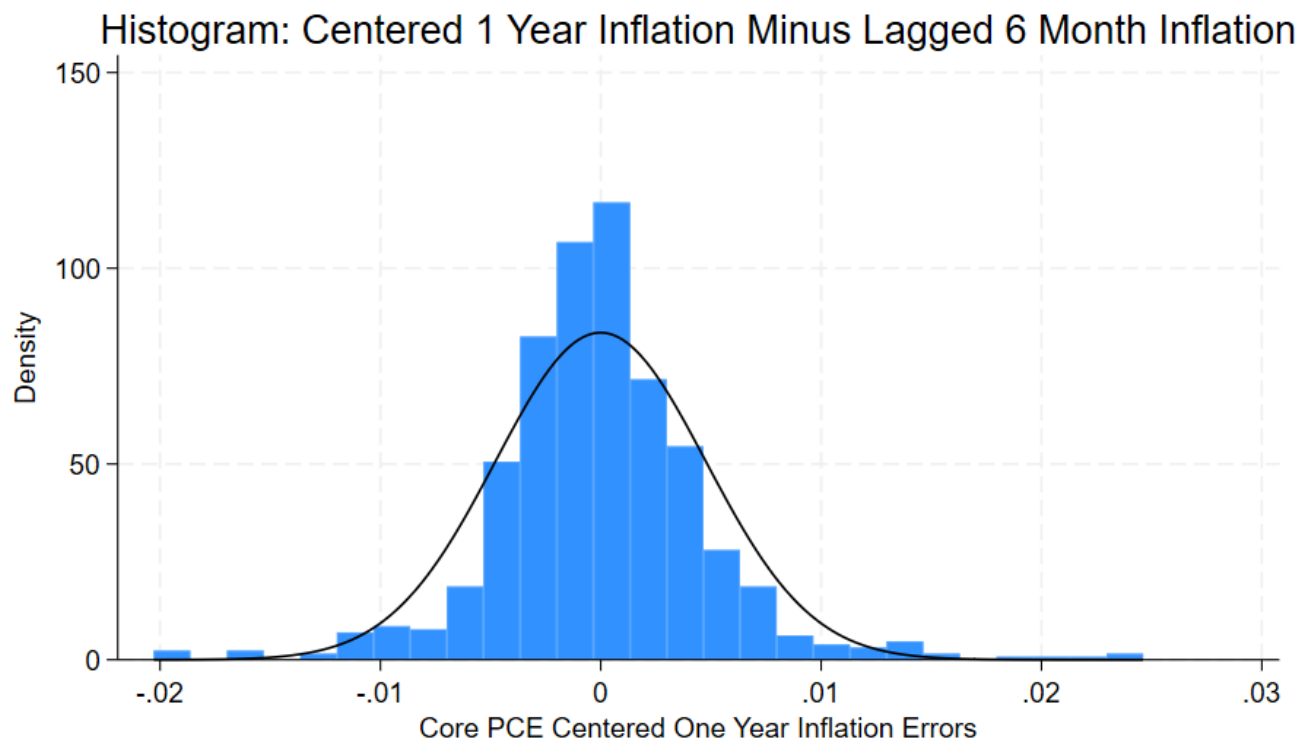
di 2*(3-2*6)/(3-6) // kurtosis = df = 6

gen double delhat6mn6ga = pL6agpcecore - L6.pcagC13mpcecore
tempvar probStultz probStugt
gen double `probStultz' = t(6,-delhat6mn6ga/.00584131)
gen double probStu6gtz58 = 1-`probStultz'
note probStu6gtz58: (1-probStultz), where probStultz = t(6,-delhat6mn6ga/.00584131). Probability
that delhat6mn6ga>0, assuming a t-dist w/ 6 df (kurt=6) and delhat standardized with
sigma=.00584131, the sd of est error for RVIE==1 & year<1995 aka 1965m1-1993m12. Created in
OptLag`vers'.do.

* For kurtosis = 6.4049, calc df
di 2*(3-2*6.4049)/(3-6.4049) // df = 5.7621663

tempvar probStultz probStugt
gen double `probStultz' = t(5.7621663,-delhat6mn6ga/.0055983)
gen double probStu6gtz55 = 1-`probStultz'
note probStu6gtz55: (1-probStultz), where probStultz = t(5.7621663,-delhat6mn6ga/.0055983).
Probability that delhat6mn6ga>0, assuming a t-dist w/ 5.7621663 df (kurt=6.4049) and delhat
standardized with sigma=.0055983, the sd of est error for 1959m7-1991m9. Created in
OptLag`vers'.do.
```

Histogram of Errors:



Section 8

References

Agresti, A., and B. A. Coull. 1998. Approximate is better than “exact” for interval estimation of binomial proportions. *American Statistician* 52: 119–126. <https://doi.org/10.1080/00031305.1998.10480550>.

Brown, L. D., T. T. Cai, and A. DasGupta. 2001. Interval estimation for a binomial proportion. *Statistical Science* 16: 101–133. <https://doi.org/10.1214/ss/1009213286>.

Clopper, C. J., and E. S. Pearson. 1934. The use of confidence or fiducial limits illustrated in the case of the binomial. *Biometrika* 26: 404–413. <https://doi.org/10.1093/biomet/26.4.404>.